Estimation and Analysis of Output Gap: An Application of Structural Vector Autoregression and Hodrick-Prescott-Fmethods

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ABSTRACT

In this study we examined output gap in the Iranian economy. The main question of the study is that how much is seasonal output gap in Iranian Economy and which factor affects gap variation. The other question is that whether using HP-F as a statistical based method for estimating output gap, provide different result than using SVAR as theory based method. Accordingly the aim of study is to estimate potential output and thus output gap using two method and analysis of the result. We used two methods (Hodrick-Prescott Filter and SVAR) to estimate quarterly output gap for the period 1988:1-2008:4. The results pointed out that the estimation is not sensitive to the method and there is a close relation between oil revenue and output gap. In the period of 1998:3-1999:3, when oil price reduced to $11.45 per barrel, Iranian economy faced with a recession and it affected on output gap with a lag. Output gap increased from 34818 in 2004:1-76782 million dollars in 2008: 4. The comparison of estimated output gap and changes of oil price in different periods point out the positive relation. According to the estimations of output gap, output gap in the Iranian Economy has intense fluctuation due to the effects of oil proceeds fluctuations. In some years, actual output is more than potential output, that is, output gap is positive and so this situation can be an important reason for inflation in that period and policy maker must do plans and policies for control of inflation and in some years, actual output is less than potential output and this means output gap was negative. This situation is a reason for unemployment in these years and therefore policy makers must do expansionary policies.

Keywords: Output Gap, Structural Var, Hodrick-Prescott Filter, Iran JEL: C13, C22, E1, E321, Iranian Economy, Expansionary Policies, Policy Makers, Intense Fluctuation

1. INTRODUCTION

The gap between actual and potential output is a key variable. While output in excess of potential, leads to higher inflation, sustained disinflation requires output to fall below potential.

If reliable, they would provide an important guide for policymakers in determining whether developments in the real economy are consistent with the maintenance of price stability. They could also assist policymakers in determining whether spending decisions and tax settings are consistent with the output gap. There are a number of reasons why central banks and governments and more generally the private sector, require accurate measures of the degree of spare capacity in the economy and the economy's long-term sustainable growth rate.

The Central Bank's need for economy-wide estimates of output gap reflects its statutory responsibility to maintain stability in the general level of prices. If the Central Bank mismeasures the level of output gap and calculates a positive output gap demand exceeds sustainable supply when in fact the gap is negative (sustainable supply exceeds demand), the consequent monetary policy tightening will tend instead

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to amplify the business cycle, offsetting some of the efficiency and allocative gains that a stable price environment may be expected to deliver. Similarly, if the Central Bank observes a rise in the rate of output growth, it needs to distinguish whether this is purely a positive demand shock and therefore possibly inflationary or simply the result of the economy responding to faster growth in the capacity of the economy to supply and therefore not inflationary. Likewise, if the Central Bank observes a fall in actual output growth, it needs to distinguish whether this is the result of a negative demand shock (and therefore possibly deflationary or a negative supply shock and therefore possibly inflationary. The concept of output gap is also important from the perspective of the Government. In the short term, an assessment of the degree of excess demand or excess capacity in the economy will influence the fiscal policy stance. As with monetary policy, an over-expansionary stance at a time when excess demand conditions are evident or an over-contractionary stance when the spare resources are plentiful will tend to amplify the business cycle. In the medium term, views regarding the economy's sustainable growth potential and thus tax base are also required to guide the formulation of the Government's fiscal strategy. This is particularly important where Government to conduct fiscal policy in a transparent and prudent manner and to achieve specific reductions in the net public debt to GDP ratio.

1.1. Literature Review

Willems (2010) estimated Output Gap and separated trend from cycle via Bayesian estimation of a New Keynesian model, augmented with an unobserved components model for output. The resulting estimate compared with popular proxies used in the literature. It turns out that the model-based approach may have important advantages for the conduct of monetary policy.

BJ (2003) estimated the Output Gap as an Indicator for the Pressure on Price Change. The output gap for 2001 was estimated as-3 to-4%, this wide output gap may be thought of as the fundamental backdrop to the continuing gradual decline in current prices.

Tommaso *et al*. (2002) estimated output gaps for Luxembourg. This study reviews several of the many alternative methods of estimating output gaps and applies six of these to annual data for Luxembourg. The sign of the output gap on the different measures seems to be systematically related, suggesting that the methods are at least measuring a related concept.

Chagny and Dopke (2001) examined Output Gap in the Euro-Zone and provided estimates of output gap in Euro Zone by using different methods including Structural VAR. the results show the methods imply different turning points and the estimated level of the output gap differs greatly.

Slevin (2001) xamined potential output and output gap in Ireland. This study estimates potential output using a number of statistical trend methods and a Cobb Douglas production function.

Cerra and Saxena (2000) examined Output Gaps with Alternative Methods including Structural VAR for Sweden. The paper reviews a number of different methods that can be used to estimate potential output and output gap. The estimates show that output gap was between-5.5 and 0.2% in period, from 1997-1998.

Gounder and Morling (2000) measured potential output in Fiji. The paper reviews four methods that can be used to estimate potential output and the output gap, including linear trends, Hodrick-Prescott (HP) filters, aggregate production functions and structural vector autoregressions. The results suggest that the output gap is measured very imprecisely in Fiji

Wedekind and Milinski (2000) estimated potential output for New Zealand by a structural VAR approach. A measure of potential output is obtained using a Structural Vector Autoregression (SVAR) methodology.

Gordon (1998) reviews five methods of estimating it for Australian GDP data, including linear time trends, Hodrick-Prescott (HP) filter trends, multivariate HP filter trends, unobservable components models and a production function model. Estimates of the gap vary with the method used and are sensitive to changes in model specification and sample period.

Israel *et al*. (1995) examined potential output by the structural vector autoregression in the Mexican economy. They find that world oil shocks have been an important source of both actual and potential output fluctuations over a sample period extending from 1965-1994.

2. MATERIALS AND METHODS

We examined output gap by two methods; Hodrick Prescott filter and SVAR. At first because Iran is oil exporting country, then we used non-oil GDP rather than GDP. For seasonal data since in quarterly data, there are seasonal fluctuations, we applied seasonal adjustment with ratio to moving average method in order to reform trend without seasonal fluctuations on data. Second we used a logarithmic transformation on quarterly GDP to obtain a more homogeneous variance of a series and avoid numerical instability.

Table 1. Unit Root Test on quarterly LGDP and first difference of it on the basis of augmented dicky -fuller and phillips perron tests

| | | Unit root test on quarterly LGDP | | | |
| | | Test critical values | | | |
	Variable	Test statistic	1%	5%	10%
Augmented	Trend and intercept	-1.3315	-4.0742	-3.4652	-3.1589
Dickey-fuller	Intercept	0.7993	-3.5121	-2.8972	-2.5855
Phillips perron	Trend and intercept	-0.5753	-4.0713	-3.4639	-3.1581
	Intercept	1.0771	-3.5101	-2.8963	-2.5851
Unit root test on first difference of quarterly LGDP					
Augmented	Trend and intercept	-4.469	-4.0819	-3.4688	-3.161
Dickey-fuller	Intercept	-3.9616	-3.5176	-2.8996	-2.5868

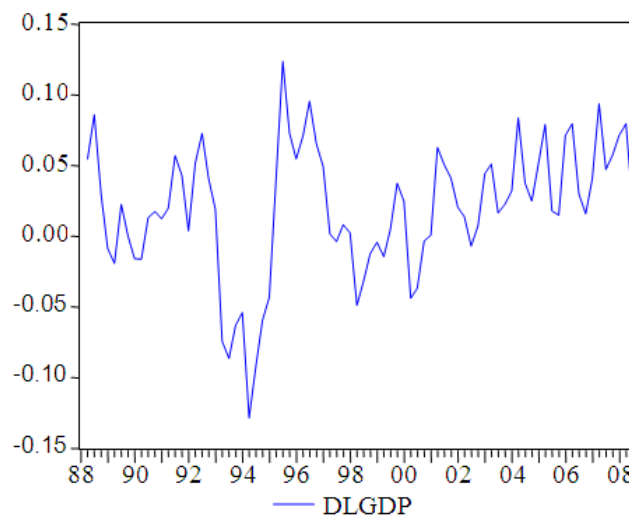**Fig. 1.** Quarterly LGDP Trend (million Dollars)

Fig. 2. First difference of quarterly LGDP Trend

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Third, we used unit root test on the logarithm of GDP (quarterly GDP data) in order to resolve spurious regression and other problems with non stationary time series.

2.1. Unit Root Test

For unit root tests, we applied Augmented Dicky-Fuller Test and Phillips Perron Test. The results show that logarithm of gross potential output is not stationary in level.

Table 1 shows the results of the augmented Dickey-Fuller and Phillips and Perron tests of the null hypothesis of non stationary GDP. The results are unambiguous and clearly support the hypothesis that the time series are stationary in first differences.

The results of Augmented Dickey-fuller test on quarterly LGDP illustrated in **Table 1**. In this case, the test statistic is in critical values whether there is a constant and/or trend included. Therefore, we conclude that quarterly LGDP is non stationary. Appropriate lag length of dependent variable can be obtained from Akaike Information (AIC), Schwartz-Baysian (SBC) and Hannan-Quinn (HQC) Criteria.

Results of Phillips Perron Unit Root Test on LGDP and critical values in different significant levels have summarized in the table. This result is similar to Augmented Dickey Fuller test results. LGDP is not stationary on its level. On the basis of Phillips Perron Unit Root Test, Null Hypothesis of LGDP has a Unit Root doesn't reject. In other words LGDP has unit root and its fluctuations around time trend is not stationary.

Therefore we try on difference of LGDP for gain of a stationary time series. After testing by unit root tests we fine that LGDP is I (1). This means that D (LGDP)-first difference of LGDP-is stationary. Augmented Dickey-fuller test on first difference of quarterly LGDP illustrated in **Table 1**. The test statistic is more than critical values whether there is a constant and/or trend included. Therefore DLGDP is stationary.

We can also find that LGDP is not stationary on level graphically. A visual plot of the data is usually the first step in the analysis of any time series. Any time series data can be thought of as being generated by a stochastic or random process and a concrete set of data, such as that shown in **Fig. 1** can be regarded as non stationary because these time series visually, at least, their mean, variance and auto covariance do not seem to be time-invariant.

The first impression that we get from the time series plotted in the following figure is that it all seems to be trending upward, although the trend is not smooth. The Quarterly GDP is in fact non stationary time series.

Figure 2 shows First-Differenced Quarterly GDP in Iran. Compared with the GDP series (Quarterly Data) given in **Fig. 1**, the differenced GDP series shown in **Fig. 2** does not show any trend. Therefore D (LGDP) in stationary. Because of D(LGDP) is stationary, as noted, it is an I(0) stochastic process, which means GDP itself is an I(1) time series and essentially it is a random walk.

2.2. Hodrick-Prescott Filter

The Hodrick-Prescott filter was created with the assumption that unobserved shocks to trend output occur all the time. The HP filter is a technique to distinguish output's long-term trend from its short-term business cycle variation. Applied economists adapted the filter by identifying the long-term trend as potential output.

The Hodrick-Prescott filter, is based on the assumption that a given time series y_t is the sum of a trend or growth component g_t and a cyclical component c_t Equation (1):

$$y_t = g_t + c_t \text{ For } t = 1, \ldots, T \tag{1}$$

According to Hodrick and Prescott (1997), "our prior knowledge is that the growth component varies 'smoothly' over time," where the measure of smoothness of the $\{g_t\}$ path is chosen to be the sum of the squares of its second difference. The cyclical component c_t represents deviations from g_t and over long time periods their average is assumed to be near zero. The growth component g_t is extracted by minimizing the following loss function Equation (2):

$$\text{Min} \left\{ \sum_{t=1}^{T} C_T^2 \lambda \sum_{t=1}^{T} C_T^2 \left[\left(g_t - g_{t-1} \right) - \left(g_{t-1} g_{t-2} \right) \right]^2 \right\} \tag{2}$$

The first sum represents the penalty for deviations of the observed series from the trend growth series $c_t = y_t - g_t$, while the second sum represents the penalty for sharp changes in the trend growth component.

We can define the filter as follows. If y denotes real GDP, the filter is defined as Equation (3):

$$\min \sum_{t=1}^{T} \left(y_t - y_t^* \right)^2 + \lambda \sum_{t=2}^{T-1} \left[\left(y_t^* - y_t^* \right) - \left(y_t - y_{t-1}^* \right) \right]^2 \tag{3}$$

With y^* as the smooth component which gives the estimate of potential GDP in this context. Broadly speaking the procedure contains two commands: (i) minimize the distance between the actual and the trend value of the time series and (ii) minimize the change of the trend value (Bhundia and Arora, 2003).
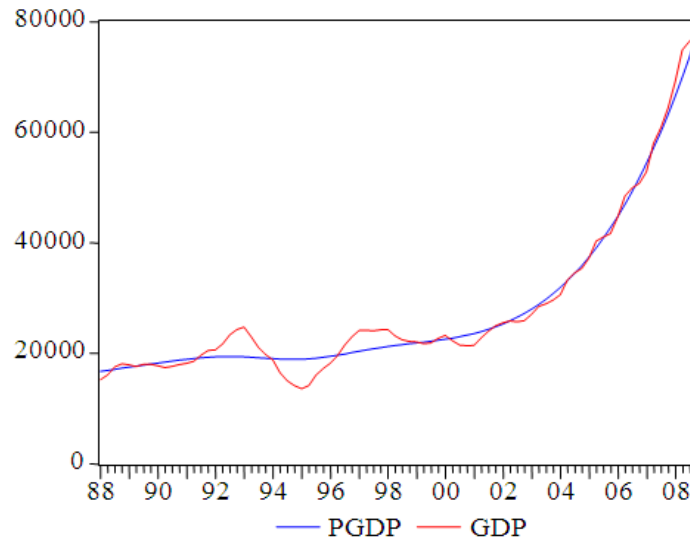
Fig. 3. Quarterly estimated potential output by Hodrick Prescott filter
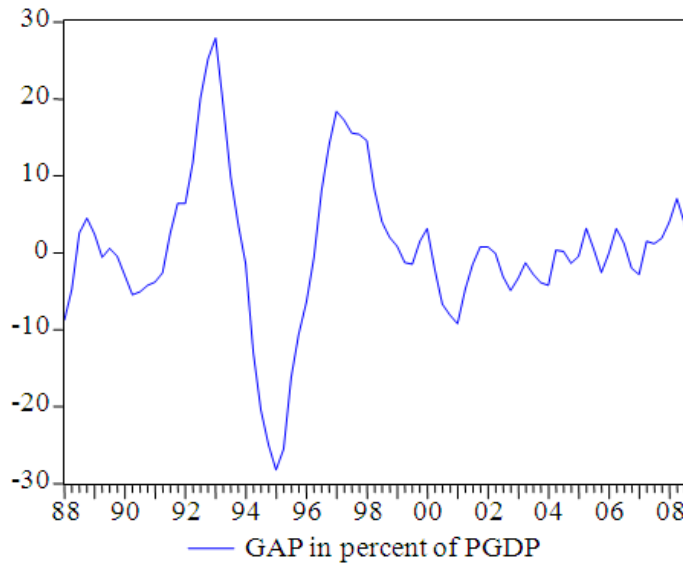

Fig. 4. Quarterly estimated output gap in percent of potential output by

Two arguments commonly made in favor of HP filter are that it extracts the relevant business-cycle frequencies of the spectrum and that it closely approximates the cyclical component implied by reasonable time series models of output (Burnside, 1998).

The filter involves the smoothing parameter λ which penalizes the acceleration in the trend component relative to the business cycle component. Researchers typically set λ = 1600 when working with quarterly data.

Most studies use the standard value of 1600 for the smoothing parameter involved in the HP-filter at the quarterly frequency as Hodrick and Prescott (1997) favored the choice of λ = 1600 (Ravn and Uhlig, 1997).

Therefore as we used quarterly data, the smoothing parameter should be 1600.

Figure 3 and 4 show potential output and output gap using Hodrick Prescott Filtering Methodology respectively.
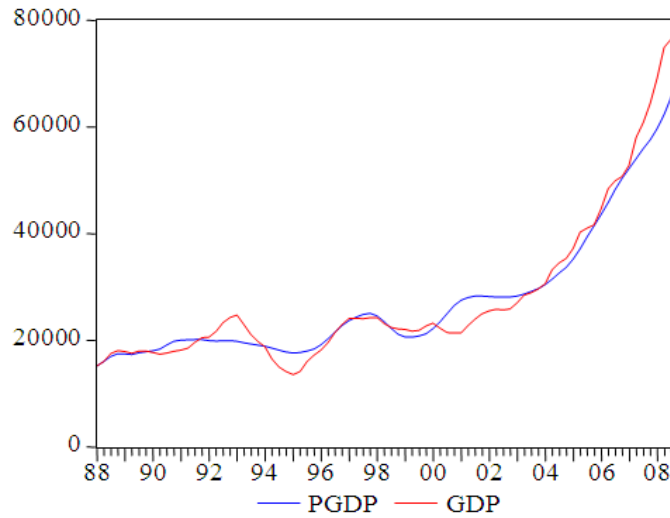
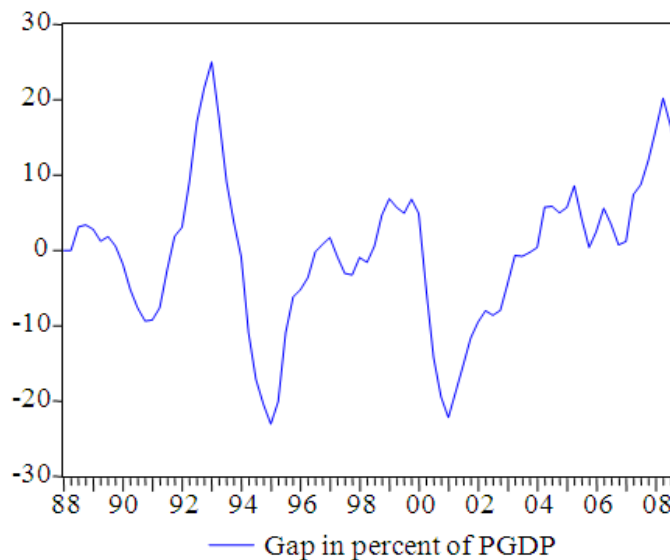**Fig. 5.** Quarterly estimated potential output by SVAR

Fig. 6. Estimated Output gap in percent of potential output by SVAR method

2.3. SVAR Method

Structural vector autoregressive is a statistical approach which estimates a vector of variables, including the change in inflation and output, as a function of those same variables.

All variables are treated as endogenous and written as a linear combination of lagged values of itself and of the other variables in the system.

Mechanical filters, such as the Hodrick and Prescott (1997) filter, is technique that extract a trend measure from actual output series. However, these filters as we said in previous section have been criticized such as the HP filter with (nearly) integrated data can induce spurious cyclicality and also the HP filters do not accurately decompose time series into their trend and cyclical components when the data have the typical spectral Granger shape. The typical Granger shape is characteristic of nearly all macroeconomic time series. Moreover, this filter shows instability of estimates near the end of the sample period.

Therefore we say SVAR advantages by following cases: First, Univariate and multivariate filters often assume that the trend component in output can be characterized as a random walk, an assumption that is not maintained in the SVAR approach.

Second, unlike some methods (such as trend-based methods, filter-based methods and the Beveridge-Nelson decomposition) the SVAR approach can be given an economic interpretation. For example, we can interpret fluctuations in potential output as being caused by certain types of shocks (Demand and supply shocks) whereas the other methods cannot. Third, contrary to other methods, such as those based on the Hodrick-Prescott filter, the VAR method does not require the imposition of an arbitrary smoothing parameter. Fourth, it theoretically overcomes the "end point problem" inherent in two sided filters (Blanchard and Quah, 1989; King et al., 1991).

We examine output gap through an approach derived from the structural vector autoregression (SVAR) methodology developed by Shapiro and Watson (1988); Blanchard and Quah (1989) and King et al. (1991). This methodology involves the estimation of a Vector Autoregression (VAR) model for the particular economy under study. We then identify different variables by making long-term assumptions based on macroeconomic theory. In order to distinguish among various sources of output fluctuation, we apply a variant of the structural VAR methodology to an autoregressive system composed of five variables, each follow a stationary stochastic process. It is assumed that the private consumption, the rate of growth of output, the monetary aggregate are endogenous and the price of oil

and a dummy variable as defined in phase trend method are exogenous to the Iranian economy in the long term.

While unit-root tests suggest that the model variables are nonstationary in levels, it is still possible that a stationary linear combination of the variables could be found. In such a case, a vector error-correction model would have to be estimated, since estimating a VAR in first differences would remove important information about the behaviour of the variables that is contained in the common trend. We used the method proposed by Johansen (1988) and applied by Johansen and Juselius (1990) to test for cointegration between the four variables. The results of the tests support the hypothesis that there is cointegrating relationship between the variables considered in this study. Note that we also applied the single-equation procedure suggested by Engle and Granger (1987) and find evidence of cointegration using that approach either. Therefore the tests reject the null hypothesis of no cointegration. Consequently, we assume that the series in the model are cointegrated and that it is appropriate to estimate the VAR models in levels (of the logarithms). We estimated Multivariate VAR model with three endogenous variables and two exogenous ones. **Figure 5 and 6** illustrate potential output and output gap using SVAR Method respectively.

3. RESULTS

In **Table 2** numerical results of potential output illustrated by two methods.
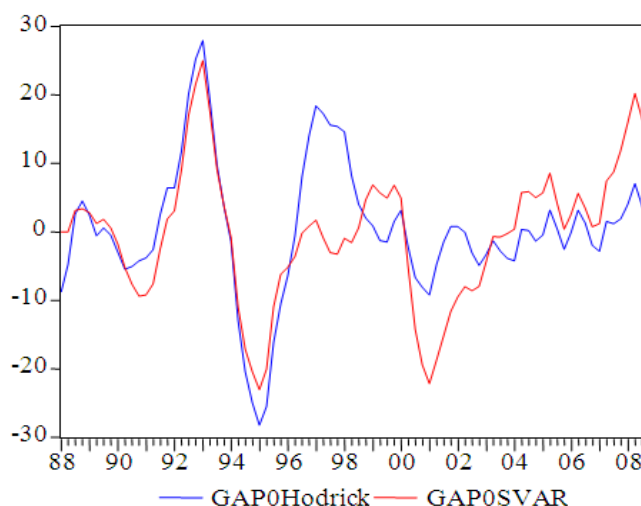

Fig. 7. Quarterly estimated output gap in percent of potential output by two methods

Table 2. Quarterly numerical results of estimated potential output (million dollars)

Range	GDP*	HP	SVAR	Range	GDP*	HP	SVAR
1988:01:00	15233	16696	-	1999:01:00	22058	21869	20646
1988:02:00	16088	16893	-	1999:02:00	21743	22024	20573
1988:03:00	17533	17092	17006	1999:03:00	21853	22185	20831
1988:04:00	18065	17290	17475	1999:04:00	22691	22357	21250
1989:01:00	17912	17487	17427	2000:01:00	23256	22546	22170
1989:02:00	17573	17681	17366	2000:02:00	22262	22756	23453
1989:03:00	17975	17872	17653	2000:03:00	21458	22995	25017
1989:04:00	17977	18060	17874	2000:04:00	21385	23268	26537
1990:01:00	17697	18243	18024	2001:01:00	21404	23580	27511
1990:02:00	17414	18420	18377	2001:02:00	22792	23937	28046
1990:03:00	17647	18592	19114	2001:03:00	23977	24342	28285
1990:04:00	17959	18755	19818	2001:04:00	24982	24798	28280
1991:01:00	18184	18907	20036	2002:01:00	25504	25309	28189
1991:02:00	18547	19046	20079	2002:02:00	25861	25878	28107
1991:03:00	19639	19167	20155	2002:03:00	25682	26511	28101
1991:04:00	20499	19266	20123	2002:04:00	25873	27212	28115
1992:01:00	20578	19341	19967	2003:01:00	27046	27986	28307
1992:02:00	21672	19388	19868	2003:02:00	28465	28839	28666
1992:03:00	23308	19405	19916	2003:03:00	28943	29775	29175
1992:04:00	24277	19391	19970	2003:04:00	29604	30799	29684
1993:01:00	24748	19348	19798	2004:01:00	30568	31917	30455
1993:02:00	22975	19282	19552	2004:02:00	33241	33133	31441
1993:03:00	21074	19199	19318	2004:03:00	34513	34452	32608
1993:04:00	19787	19109	19087	2004:04:00	35383	35878	33702
1994:01:00	18750	19023	18907	2005:01:00	37236	37415	35227
1994:02:00	16491	18951	18519	2005:02:00	40306	39070	37130
1994:03:00	15046	18905	18155	2005:03:00	41033	40847	39376
1994:04:00	14177	18892	17822	2005:04:00	41651	42750	41495
1995:01:00	13579	18919	17648	2006:01:00	44727	44787	43638
1995:02:00	14147	18990	17699	2006:02:00	48438	46962	45872
1995:03:00	16010	19104	17982	2006:03:00	49916	49281	48267
1995:04:00	17224	19256	18361	2006:04:00	50719	51750	50359
1996:01:00	18192	19442	19188	2007:01:00	52822	54373	52210
1996:02:00	19532	19654	20268	2007:02:00	58011	57159	54016
1996:03:00	21491	19883	21534	2007:03:00	60818	60109	55934
1996:04:00	22946	20121	22762	2007:04:00	64414	63228	57574
1997:01:00	24105	20360	23703	2008:01:00	69181	66521	59679
1997:02:00	24147	20593	24353	2008:02:00	74912	69991	62332
1997:03:00	24058	20815	24815	2008:03:00	76451	73644	65581
1997:04:00	24255	21022	25074	2008:04:00	77178	77486	68693
1998:01:00	24310	21215	24545				
1998:02:00	23152	21392	23526				
1998:03:00	22424	21558	22300				
1998:04:00	22150	21715	21163				

*Non-oil GDP after seasonal adjustment with ratio to moving average method

In **Fig. 7** it was shown the results of estimated output gap.

The two measures of output gap were estimated for Iran. The results are shown in **Fig. 7 and Table 2**. Looking at **Fig. 7**, the graph shows that each method almost produces a similar result for output gap. The results pointed out that the estimation is not sensitive to the method.

4. DISCUSSION

Although results are almost near together but we have to select one method for analysis. Therefore we analyze the obtained results in economic viewpoint now by SVAR method for quarterly data because of SVAR method having the highest magnitude as we explained in previous sections.
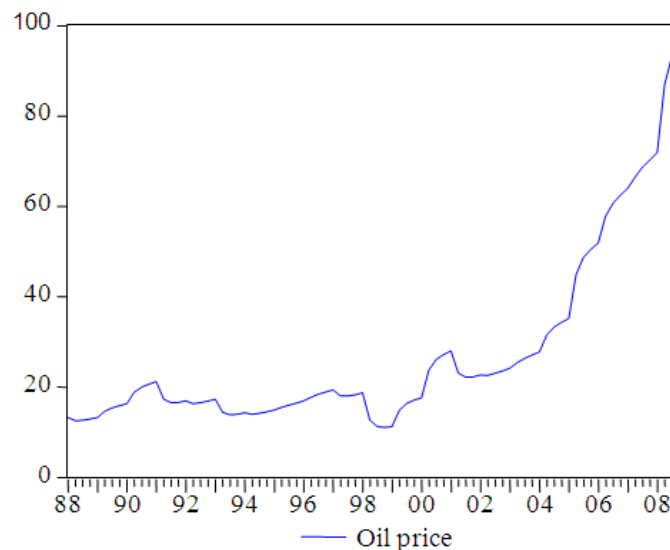
Fig. 8. Crude oil price per barrel(dollars)

The mean of output gap is-0.09% with 27% maximum and-23% minimum levels so that indicating severe fluctuations of output gap in Iranian economy in the period as depicted graphically.

This means that estimated output gap from this method fluctuates about between 27 and-23 in percent of potential output. The highest and lowest levels are related to 1993:1 and 1995:1 respectively. The nearness of the maximum and minimum of output gap (1993:1 and 1995:1) shows instability and inconstancy in output gap level in Iranian economy. Therefore the question rise for what reason the fluctuation happened?

The plausible reason is that Iranian economy has relying upon oil income strongly, thus we measure correlation of two variables, e.g., potential output and oil income. Results show strong correlation between these variables (about 97%) also this problem is applicable about potential output and liquidity so that correlation is about 98%. The high correlation indicates tightly relation between oil income and potential output.

Correlation does not necessarily imply causation in any meaningful sense of that word. While there are high positive correlation between two variables but we need to examine the causality. Therefore we test Granger Causality Test for finding of causality and its direction.

The casualityapproach to the question of whether potential output causes oil income or the latter one causes the former. According to the results, we cannot reject the hypothesis that potential output does not Granger cause oil income but we do reject the hypothesis that oil income does not Granger cause potential output.

In other words oil income Granger causes potential output and then we can state a part of potential output is the effect or the result of oil income. Therefore it appears that Granger causality runs one-way from oil income to potential output and not the other way. Consequently oil income influences on potential output in Iranian economy.

The conclusion helps us better and more precise analysis of potential output changes. As it has shown in **Fig. 8** in period 1988:1-1991:3 the oil price increased (from 12.85 $ in 1988:1-19.91 $ per barrel in 1990:4 for heavy oil) and therefore potential output increased from 20296 in 1988:3-24044 million dollars in 1991:3. Although increasing of oil price and potential output are not coordinated completely because of increasing oil price effect in economy with delay naturally.

In period of 1998:3-1999:3, there was a recession because of reducing oil price as it reduced to 11.45 $ a barrel. Iranian economy has been faced with a severe recession and it affects on output gap with delay. Reducing of oil price is an unfavorable shock because it affects on GDP, national income and ultimately output gap negatively. It is important that oil shocks have two effects on economy, favorable and unfavorable effect. In Iran economy, favorable oil shocks like the OPEC oil price increases of the 2000s raise potential output in the long run. The oil price decreases which occurred in late 1990s were unfavorable oil shocks.

This process repeats more or less all of the period especially at the end of it, whereas potential output

increased from 34818 in 2004:1-76782 million dollars in 2008:4. In other words it faces to increasing 120% from beginning to ending of period and 4.25% rate of growth seasonally whereas the oil price has increased from 33.06-91.49 dollars a barrel.

From an economic viewpoint, the effect of oil price on output gap point out that oil proceed has fundamental role in the economy so that the long run effect is a possible rising in the level of potential GDP or potential output. Thus, the long-run aggregate supply curve line may shift to the right in response to a supply shock. We see this issue in **Fig. 5 and 6** so that oil shock in 2004-2008 causes increasing potential output and its trend is uprising.

5. CONCLUSION

According to the estimations of output gap, output gap in the Iranian Economy has intense fluctuation due to the effects of oil proceeds fluctuations. Therefore there are close relationship between oil proceeds and output gap in the Iranian Economy. In some years, actual output is more than potential output, that is, output gap is positive and so this situation can be an important reason for inflation in that period and policy maker must do plans and policies for control of inflation and in some years, actual output is less than potential output and this means output gap was negative. This situation is a reason for unemployment in these years and therefore policy makers must do expansionary policies. The comparison of estimated output gap and changes of oil price in different periods point out the positive relation. The estimation is also insensitive to the method of estimation and provides almost identical results.

We suggest potential output and therefore output gap are estimated every season. Policy makers must have more attention to potential output level and output gap. It can help to perform better economic five years plans. This can help to reduce unemployment and inflation.

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