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Estimation under Multicollinearity: A Comparative Approach Using Monte Carlo Methods

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Abstract: Problem statement: A comparative investigation was done experimentally for 6 different Estimation Techniques of a just-identified simultaneous three-equation econometric model with three multi-collinear exogenous variables. **Approach:** The aim is to explore in depth the effects of the problems of multicollinearity and examine the sensitivity of findings to increasing sample sizes and increasing number of replications using the mean and total absolute bias statistics. **Results:** Findings revealed that the estimates are virtually identical for three estimators: LIML, 2SLS and ILS, while the performances of the other categories are not uniformly affected by the three levels of multicollinearity considered. It was also observed that while the frequency distribution of average parameter estimates was rather symmetric under the OLS, the other estimators was either negatively or positively skewed with no clear pattern. **Conclusion:** The study had established that L2ILS estimators are best for estimating parameters of data plagued by the lower open interval negative level of multicollinearity while FIML and OLS respectively rank highest for estimating parameters of data characterized by closed interval and upper categories level of multicollinearity.

Key words: Multicollinearity, Monte-Carlo, simultaneous equation, just-identification and exogenous variables

INTRODUCTION

One of the most frequently suggested solutions to the problem of multicollinearity in single equation estimation is the use of simultaneous econometric model. In the simultaneous model, the problem of multicollinearity may still exist in the individual equations. If the simultaneous equation solution to this problem is adopted there may be an intolerable rise in the size of the model with the consequent depletion of the number of exogenous variables which are usually required for policy simulation.

Since some measure of multicollinearity has to be tolerated in simultaneous econometric models, this study therefore investigates the comparative performance of six different estimation techniques namely: Ordinary Least Squares (OLS), Limited Information Maximum Likelihood (LIML), Two-Stage Least Squares (2SLS), Indirect Least Squares (ILS), Three-Stage Least Squares (3SLS) and Full Information Maximum Likelihood (FIML) under three different levels of multicollinearity between the multicollinear exogenous variables. The performances of the estimators are evaluated based on the average or mean values of parameter estimates and total absolute bias of

parameter estimates. The aim is to explore in depth the phenomena effects and examine the sensitivity of findings to increasing sample sizes and increasing number of replications Goodnight and Wallace (1969), Hoerl and Kennard (1970) and Goodnight and Wallace (1972).

Studies on estimation under multicollinearity effects of simultaneous models revealed that a high degree of multicollinearity among the explanatory variables has a disastrous effect on estimation of the coefficients, β by the OLS Fisher (1966). This method was considered by Hendry (1976), RAY, (1970) and Pleli and Tankovic (2005) as naive approach because the estimators are biased and inconsistent. They however categorized other methods as limitedinformation approach (2SLS, ILS) and full-information approach (3SLS, FIML). Adenomon and Fesojaiye (2008), Agunbiade and Osilagun (2008) merelv compared the Seemingly Unrelated Regression (SUR) with the OLS technique and confirmed the superiority of the SUR estimator to the OLS estimators. In the opinion of Ayinde (2007), where he compared OLS with some GLS estimators, he observed that with increasing replications OLS estimator is preferred in estimating all the model parameters at all levels of correlation. However, this opinion negates Pleli and

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Tankovic (2005) in which they advised an econometrician to avoid the use of naïve approach (OLS) in estimating the parameters of a system of simultaneous equations.

Framework of the model: In this study, a Monte-Carlo approach is employed for the following just-identified econometric model having three structural equations:

$$y_{1t} = \beta_{13}y_{3t} + \gamma_{11}x_{1t} + \gamma_{12}x_{2t} + u_{1t}$$

$$y_{2t} = \beta_{21}y_{1t} + \gamma_{21}x_{1t} + \gamma_{23}x_{3t} + u_{2t}$$

$$y_{3t} = \beta_{32}y_{2t} + \gamma_{32}x_{2t} + \gamma_{33}x_{3t} + u_{3t}$$
(1)

Where:

y_{1t} , y_{2t} and y_{3t}	=	Endogenous	or	jointly	dep	endent
		variables				
x_{1t} , x_{2t} and x_{3t}	=	Predetermined	(e	xogenous	or	lagged
		endogenous) v	aria	ables		

The ult, u_{2t} and u_{3t} denote stochastic disturbance terms which are assumed to be independently and identically normally distributed with zero means and finite covariance matrix, β_{13} , β_{21} , β_{32} are coefficients of endogenous variables while the γ_{11} , γ_{12} , γ_{21} , γ_{23} , γ_{32} , γ_{33} are the coefficients of predetermined variables making nine structural parameters for the model.

Express the model (1) in matrix form yields:

$$y = x \beta + u \tag{2}$$

MATERIALS AND METHODS

The methodology employed in this study is the Monte-Carlo Approach (MCA). The Monte-Carlo method the nearest thing to a controlled is type experiment in econometrics laboratory Intrilligator al. (1996); Johnston (1984), et Agunbiade (2007), Carlin et al., (1992), Kmenta and Joseph (1963), Parker (1972), Wagnar (1958), Olayemi and Olayide (1981) and Koutsoyiannis (2008). The MCA has been applied not only to Multicollinearity effect but also to choice of alternative estimators in determining the impact of heteroscedasticity, serial correlation and other violations of basic econometric assumptions on the performance of different estimators in a given study. It is also used to solve problems on both pure and social sciences Belsley et al., (1992), Farrar and Glauber (1967), Feldstein (1973) and Mishra (2004).

In order to assemble data that will conform to the model specified, our data series are generated as follows:

- 1. We set sample sizes N at 100, 200 and 300 and replication numbers R = 200, 400 and 600 for this study. These values are arbitrary although it compares favorably with sample sizes in other similar studies
- 2. The following numerical values are arbitrarily assigned to each of the structural parameters of the model:

$$\beta_{13} = 1.8 \gamma_{11} = 0.2 \gamma_{12} = 1.2$$

$$\beta_{21} = 1.5 \gamma_{21} = 2.5 \gamma_{23} = 2.1$$

$$\beta_{32} = 0.9 \gamma_{32} = 0.4 \gamma_{33} = 3.3$$
(3)

3. Values are assigned to each of the elements of the variance-covariance matrix of the disturbance terms of the model at any given sample points:

$$\sum_{e} = \begin{pmatrix} 7.0 & 5.0 & 4.0 \\ 5.0 & 4.5 & 3.5 \\ 4.0 & 3.5 & 3.0 \end{pmatrix}$$
(4)

- 4. Values of the predetermined variables x_{1t} , x_{2t} and x_{3t} are generated from a pool of uniformly (0,1) distributed random numbers Kmenta (1971) using the Microsoft Excel package such that the correlation coefficients ρ (x_1 , x_2), $\rho(x_2$, x_3) and $\rho(x_1, x_3)$ are in the following ranges of the three levels of multicollinearity considered:
 - Relatively highly negatively correlated (ρ_{xi,xj}<-0.05) which is referred to as Lower Open Interval Negative (LON)
 - Feebly negatively or positively correlated (-0.05≤ρ_{xi,xj}≤+0.05) which is referred to as Closed Interval Negative or Positive (CNP)
 - Relatively highly positively correlated $(\rho_{xi,xj} > +0.05)$ which is termed Upper Open Interval Positive (UOP)

Consequently there are three sets of X's in each category of the multicollinearity group. We perform the correlation matrices to ascertain the usefulness of data set:

- 5. Values of the disturbance terms u_{1t} , u_{2t} and u_{3t} are specified to each sample point. A two-stage process is employed to generate these values:
 - Three sets of random normal series were generated and standardized to obtain independent series ϵ_t of random normal deviates
 - The generated series are transformed into three series of random disturbance in order to obtain covariance matrix predetermined in step (3)

above for the model. The method presented by Nagar (1969) for transformation of independent series of standard random deviates into series of random deviates with zero mean and specified variance-covariance matrices is used for this purpose. This is described below

According to Nagar (1969), since Σ is a positive definite matrix, we can decompose it by a non-singular upper triangular matrix P such that:

$$\sum_{c} \otimes I_{n} = PP'$$
 (5)

Letting
$$P = \begin{pmatrix} S_{11} & S_{12} & S_{13} \\ 0 & S_{22} & S_{23} \\ 0 & 0 & S_{33} \end{pmatrix}$$
 (6)

and Eq. 6 to:

$$\begin{pmatrix} \sigma_{11} & \sigma_{12} & \sigma_{13} \\ 0 & \sigma_{22} & \sigma_{23} \\ 0 & 0 & \sigma_{33} \end{pmatrix}$$

where, $\sigma_{ij} = \sigma_{ji}$, $i \neq j$ it can be shown that:

$$\begin{split} \mathbf{S}_{33} &= +\sqrt{\sigma_{33}} \\ \mathbf{S}_{23} &= \sigma_{33} / \sqrt{\sigma_{33}} \\ \mathbf{S}_{22} &= \left(\sigma_{12}\sigma_{23} / \sigma_{33}\right)^{1/2} \\ \mathbf{S}_{13} &= \sigma_{31} / \sqrt{\sigma_{33}} \\ \mathbf{S}_{12} &= \left(\sigma_{12} - \mathbf{S}_{23}\mathbf{S}_{13}\right) / \mathbf{S}_{22} \\ \mathbf{S}_{11} &= +\sqrt{\left(\sigma_{11} - \mathbf{S}_{12}^2 - \mathbf{S}_{13}^2\right)} \end{split}$$
(7)

The three random disturbance series are thus formed using:

$$\mathbf{U}_{t} = \begin{pmatrix} \mathbf{u}_{1t} \\ \mathbf{u}_{2t} \\ \mathbf{u}_{3t} \end{pmatrix} = \mathbf{P} \boldsymbol{\varepsilon}_{t} = \begin{pmatrix} \mathbf{S}_{11} & \mathbf{S}_{12} & \mathbf{S}_{13} \\ \mathbf{0} & \mathbf{S}_{22} & \mathbf{S}_{23} \\ \mathbf{0} & \mathbf{0} & \mathbf{S}_{33} \end{pmatrix} \begin{pmatrix} \boldsymbol{\varepsilon}_{1t} \\ \boldsymbol{\varepsilon}_{2t} \\ \boldsymbol{\varepsilon}_{3t} \end{pmatrix}$$
(8)

Hence:

$$U_{1t} = S_{11}\varepsilon_{1t} + S_{12}\varepsilon_{2t} + S_{13}\varepsilon_{3t}$$

$$U_{2t} = S_{22}\varepsilon_{2t} + S_{23}\varepsilon_{3t}$$

$$U_{3t} = S_{13}\varepsilon_{3t}$$
(9)

6. The endogenous variables are then generated from the values already obtained for the X's (step 4) and the U's (step 5) and the values assigned to the structural parameters (step (2)). This is most conveniently done using the reduced form model derived as follows:

Using our three-equation model:

$$y_{1t} = \beta_{13}y_{3t} + \gamma_{11}x_{1t} + \gamma_{12}x_{2t} + u_{1t}$$

$$y_{2t} = \beta_{21}y_{1t} + \gamma_{21}x_{1t} + \gamma_{23}x_{3t} + u_{2t}$$

$$y_{3t} = \beta_{32}y_{2t} + \gamma_{32}x_{2t} + \gamma_{33}x_{3t} + u_{3t}$$

Rearranging the model we have:

$$\begin{aligned} y_{1t} + 0y_{2t} - \beta_{13}\gamma_{3t} - \gamma_{11}x_{1t} - \gamma_{12}x_{2t} + 0x_{3t} = U_{1t} \\ -\beta_{21}y_{1t} + y_{2t} + 0y_{3t} - \gamma_{21}x_{1t} + 0x_{2t} - \gamma_{23}x_{3t} = U_{2t} \\ 0y_{1t} - \beta_{32}y_{2t} + \gamma_{3t} - 0x_{1t} - \gamma_{32}x_{2t} - \gamma_{33}x_{3t} = U_{3t} \end{aligned}$$

This can be written as:

$$\beta \mathbf{y}_t + \Gamma \mathbf{X}_t = \mathbf{U}_t \tag{10}$$

Where:

$$\beta = \begin{pmatrix} 1 & 0 & -\beta_{13} \\ -\beta_{21} & 1 & 0 \\ 0 & -\beta_{32} & 1 \end{pmatrix}$$
$$\Gamma = \begin{pmatrix} -\gamma_{11} & -\gamma_{12} & 0 \\ -\gamma_{21} & 0 & -\gamma_{23} \\ 0 & -\gamma_{32} & -\gamma_{33} \end{pmatrix}$$
$$y_{t} = \begin{pmatrix} y_{1t} \\ y_{2t} \\ y_{3t} \end{pmatrix}$$
$$x_{t} = \begin{pmatrix} x_{1t} \\ x_{2t} \\ x_{3t} \end{pmatrix}$$
$$U_{t} = \begin{pmatrix} U_{1t} \\ U_{2t} \\ U_{3t} \end{pmatrix}$$

Rewriting Eq. 10, to make y_t the subject of the relations, we have:

$$\mathbf{y}_{t} = -\beta^{-1} \Gamma \mathbf{X}_{t} + \beta^{-1} \mathbf{U}_{t} \tag{11}$$

Where:

$$\beta^{-1} = \frac{1}{1 - \beta_{13}\beta_{32}\beta_{21}} \begin{pmatrix} 1 & \beta_{13}\beta_{32} & \beta_{13} \\ \beta_{21} & 1 & \beta_{21}\beta_{13} \\ \beta_{32}\beta_{21} & \beta_{32} & 1 \end{pmatrix}$$

So:

$$\begin{split} y_{t} = & \frac{-1}{1 - \beta_{13}\beta_{32}\beta_{21}} \begin{pmatrix} 1 & \beta_{13}\beta_{32} & \beta_{13} \\ \beta_{21} & 1 & \beta_{21}\beta_{13} \\ \beta_{32}\beta_{21} & \beta_{32} & 1 \end{pmatrix} \begin{pmatrix} -\gamma_{11} & -\gamma_{12} & 0 \\ -\gamma_{21} & 0 & -\gamma_{23} \\ 0 & -\gamma_{32} & -\gamma_{33} \end{pmatrix} \\ & \begin{pmatrix} x_{1t} \\ x_{2t} \\ x_{3t} \end{pmatrix} + \frac{-1}{1 - \beta_{13}\beta_{32}\beta_{21}} \begin{pmatrix} 1 & \beta_{13}\beta_{32} & \beta_{13} \\ \beta_{21} & 1 & \beta_{21}\beta_{13} \\ \beta_{32}\beta_{21} & \beta_{32} & 1 \end{pmatrix} \begin{pmatrix} U_{1t} \\ U_{2t} \\ U_{3t} \end{pmatrix} \end{split}$$

Alternatively, Eq. 11 can be written in terms of its reduced form parameters:

$$\mathbf{y}_{t} = \Pi \mathbf{X}_{t} + \mathbf{V}_{t} \tag{12}$$

where, Π is the reduced form of parameters defined as:

 $\Pi = -\beta^{-1}\Gamma$

and:

$$\mathbf{V}_{t} = \boldsymbol{\beta}^{-1} \mathbf{u}_{t} \tag{13}$$

So:

$$\mathbf{y}_{t} = \begin{pmatrix} \Pi_{11} & \Pi_{12} & \Pi_{13} \\ \Pi_{21} & \Pi_{22} & \Pi_{23} \\ \Pi_{31} & \Pi_{32} & \Pi_{33} \end{pmatrix} \begin{pmatrix} \mathbf{x}_{1t} \\ \mathbf{x}_{2t} \\ \mathbf{x}_{3t} \end{pmatrix} + \begin{pmatrix} \mathbf{V}_{1t} \\ \mathbf{V}_{2t} \\ \mathbf{V}_{3t} \end{pmatrix}$$
(14)

Equation 14 is expressed as:

$$\begin{aligned} y_{1t} &= \Pi_{11} X_{1t} + \Pi_{12} x_{2t} + \Pi_{13} x_{3t} + V_{1t} \\ y_{2t} &= \Pi_{21} X_{1t} + \Pi_{22} x_{2t} + \Pi_{23} x_{3t} + V_{2t} \\ y_{3t} &= \Pi_{31} X_{1t} + \Pi_{32} x_{2t} + \Pi_{33} x_{3t} + V_{3t} \end{aligned} \tag{15}$$

So:

$$\begin{split} \Pi &= \frac{-1}{1 - \beta_{13}\beta_{32}\beta_{21}} \begin{pmatrix} 1 & \beta_{13}\beta_{32} & \beta_{13} \\ \beta_{21} & 1 & \beta_{21}\beta_{13} \\ \beta_{32}\beta_{21} & \beta_{32} & 1 \end{pmatrix} \begin{pmatrix} \gamma_{11} & \gamma_{12} & 0 \\ \gamma_{21} & 0 & \gamma_{23} \\ 0 & \gamma_{32} & \gamma_{33} \end{pmatrix} \\ &= \frac{1}{D} \begin{pmatrix} \gamma_{11} + \beta_{13}\beta_{32}\gamma_{21} & \gamma_{12} + \beta_{13}\gamma_{32} & \beta_{13}\beta_{32}\gamma_{23} + \beta_{13}\gamma_{33} \\ \beta_{21}\gamma_{11} + \gamma_{21} & \beta_{21}\gamma_{12} + \beta_{21}\beta_{13}\gamma_{32} & \gamma_{23} + \beta_{21}\beta_{13}\gamma_{33} \\ \beta_{32}\beta_{21}\gamma_{11} + \beta_{32}\gamma_{21} & \beta_{32}\beta_{21}\gamma_{12} + \gamma_{32} & \beta_{32}\beta_{21}\gamma_{23} + \gamma_{33} \end{pmatrix} \end{split}$$

Thus we have:

$$\mathbf{V}_{t} = \beta^{-1}\mathbf{u}_{t} = \frac{1}{D} \begin{pmatrix} \mathbf{U}_{1t} & +\beta_{13}\beta_{32}\mathbf{U}_{2t} & +\beta_{13}\mathbf{U}_{3t} \\ \beta_{21}\mathbf{U}_{1t} & +\mathbf{U}_{2t} & +\beta_{21}\beta_{13}\mathbf{U}_{3t} \\ \beta_{32}\beta_{21}\mathbf{U}_{1t} & +\beta_{32}\mathbf{U}_{2t} & +\mathbf{U}_{3t} \end{pmatrix}$$
(16)

Thus we have:

$$\begin{split} \mathbf{y}_{1t} &= \left(\frac{\gamma_{11} + \beta_{13}\beta_{32}\gamma_{21}}{D}\right) \mathbf{X}_{1t} + \left(\frac{\gamma_{12} + \beta_{13}\gamma_{32}}{D}\right) \\ &\mathbf{X}_{2t} + \left(\frac{\beta_{13}\beta_{32}\gamma_{23} + \beta_{13}\gamma_{33}}{D}\right) \\ &\mathbf{X}_{3t} + \left(\frac{U_{1t} + \beta_{13}\beta_{32}U_{2t} + \beta_{13}U_{3t}}{D}\right) \\ &\mathbf{y}_{2t} = \left(\frac{\beta_{21}\gamma_{11} + \gamma_{21}}{D}\right) \mathbf{X}_{1t} + \left(\frac{\beta_{21}\gamma_{12} + \beta_{21}\beta_{13}\gamma_{32}}{D}\right) \\ &\mathbf{X}_{2t} + \left(\frac{\gamma_{23} + \beta_{21}\beta_{13}\gamma_{33}}{D}\right) \\ &\mathbf{X}_{3t} \left(\frac{\beta_{21}U_{1t} + U_{2t} + \beta_{21}\beta_{13}U_{3t}}{D}\right) \\ &\mathbf{y}_{3t} = \left(\frac{\beta_{32}\beta_{21}\gamma_{11} + \beta_{32}\gamma_{21}}{D}\right) \mathbf{X}_{1t} + \left(\frac{\beta_{32}\beta_{21}\gamma_{12} + \gamma_{32}}{D}\right) \\ &\mathbf{X}_{2t} + \left(\frac{\beta_{32}\beta_{21}U_{1t} + \beta_{32}U_{2t} + U_{3t}}{D}\right) \\ &\mathbf{X}_{3t} + \left(\frac{\beta_{32}\beta_{21}U_{1t} + \beta_{32}U_{2t} + U_{3t}}{D}\right) \end{split}$$

Where:

 $D = -(1 - \beta_{13}\beta_{32}\beta_{21})^{-1}$

The Eq. 17 is used to determine the values of the endogenous variables at each sample point.

- The final stage of this experiment is the estimation of the structural parameters with the aid of the generated data sets y_{1t}, y_{2t}, y_{3t}, x_{1t}, x_{2t} and x_{3t}. The following estimation methods are employed:
- Ordinary least squares method
- Two stage least squares method
- Limited information maximum likelihood method
- Indirect least squares method
- Three-stage least squares method
- Full-information maximum likelihood method

RESULTS AND DISCUSSION

In theory and as confirmed by Johnston (1984), when an equation is just identified, estimates of parameters obtained by 2SLS, LIML, ILS and 3SLS should be identical. However, the results obtained among the six estimation techniques used in the study revealed that the estimates are virtually identical for the three estimators: LIML, 2SLS and ILS (referred to as L2ILS). The performance of the four categories of the estimators (OLS, L2ILS, 3SLS and FIML) are not uniformly affected by the three levels of multicollinearity. For the three cases of multicollinear exogenous variables the frequency distribution of average parameter estimates under FIML, 3SLS and L2ILS was either negatively or positively skewed with no clear pattern while the distribution was rather symmetric under the OLS (Fig. 1 and 2). However, the performance of these estimators improved better as sample size increased.

A comparative performance evaluation of the four categories of estimators using the Average of parameter Estimates revealed that:

- The parameters of the three equations are underestimated in 84 present of the entire average estimates, this observation is more prevalent for Eq. 1 than any of the other two. Under estimation is more serious in respect of OLS and 3SLS
- CNP contains estimates that are closest to the true parameter values than the other two open intervals



Fig. 1: Charts of the frequency distribution of OLS estimates N = 100, R = 200 for LON, CNP and UOP



Fig. 2: Charts of the frequency distribution of L2ILS estimates N = 100, R200 for LON, CNP and UOP

- FIML is best for estimating Eq. 1 at both LON and CNP while L2ILS is best at UOP. Both FIML and OLS are best for estimating Eq. 2 while FIML is recommended for estimation of Eq. 3
- Ranking of estimators over the three multicollinearity levels indicates that L2ILS estimators are best for estimating parameters of data plagued by the lower open interval negative while FIML ranks high in the closed feebly interval. Also, OLS gave a clear lead in estimating parameters of the upper open interval positive
- The estimates of most parameters of Eq. 1 and 3 are peaked at the middle closed interval where exogenous variables are feebly correlated. The reverse is noted for parameters of Eq. 2 (trend type "v") where most estimators attain their minimum when multicollinearity level is feebly closed negative or positive interval
- No remarkable asymptotic pattern is noticed in the performance of the estimates of the parameters of each estimator
- The performance of estimators is not affected by changes in the replication numbers, that is, no evidence of sensitivity of the distribution of estimators to number of replications which appears to attest to the stability of the results obtained in this study

The following are the main findings based on the use of Total Absolute Biases (TAB) of parameters estimates:

- The estimates of the absolute biases for the estimators are relatively smaller when compared with some earlier work Oduntan (2004) where he studied two just identified equations
- The Model Total Absolute Bias (MTAB) of OLS and 3SLS estimators increased asymptotically while the estimates of MTAB do not reveal any such asymptotic behavior for L2ILS and FIML (though the sample size N = 200 appears to be the turning point)
- Model total absolute bias as expected did not reveal any sensitivity to changes in replication
- The ranking of estimators based on the Average Model Total Absolute Bias (AMTAB) and the Coefficient of Variation (CV) revealed that the four estimators rank uniformly in the following order: OLS, 3SLS, FIML and L2ILS

- As correlation levels changes from LON through CNP to UOP MTAB decreased consistently (\) falls for OLS, rose consistently (/) for L2ILS and has the minimum at the middle level (\/) for FIML. The behavior is inconclusive for 3SLS
- Expectedly, the trends ranked as follows in decreasing order of frequency; the concave "√" type the downward sloping '/' and the capital lamda '∧'
- The asymptotic distribution of 'best' estimators revealed that L2ILS are best in estimating the

parameters of LON. FIML is best in the CNP while OLS consistently remained the best estimator for positively multicollinear exogenous variables (UOP). The findings are similar when average of estimates was used

The Table 1-9 in the appendices are attached as part of the tables generated in the course of the analysis. They were used to arrive at our conclusion. Also, the charts are attached to reveal our claim of the nature of distribution.

Table 1: Po	erformance of estin	nators using a	average N = 10	0, R = 200							
		Equation 1			Equation 2			Equation 3			
Estimator	Correlation level	β ₁₃ (1.8)	γ ₁₁ (0.2)	γ ₁₂ (1.2)	β ₂₃ (1.5)	γ ₂₁ (2.5)	γ ₂₃ (2.1)	β ₃₂ (0.9)	γ ₃₂ (0.4)	γ ₃₃ (3.3)	
OLS	LON	1.2754	0.9597	0.5078	1.1877	-1.6168	-1.62220	0.6659	0.0868	-1.6067	
	CNP	1.2798	0.7404	0.6351	1.1858	-1.7576	-1.72440	0.6646	0.3619	-1.6150	
	UOP	1.2766	1.1768	0.6132	1.1780	-0.9822	-1.21120	0.6685	0.0263	-1.5004	
LIML	LON	1.6295	-0.2063	-1.3197	2.0743	-1.0381	-2.34030	0.3431	0.4827	1.0405	
	CNP	1.5990	-0.2075	-0.3553	1.5521	-3.4071	-3.37100	0.7056	0.2711	-1.7960	
	UOP	1.5354	0.5446	-0.1800	-5.7211	5.3293	25.30070	0.8182	1.5853	-3.8157	
2SLS	LON	1.6295	-0.2064	-1.8119	2.0743	-1.1068	-2.34030	0.3200	0.4827	1.1718	
	CNP	1.5990	-0.2075	-0.3553	1.5520	-3.4070	-3.37120	0.7056	0.2711	-1.7961	
	UOP	1.5354	0.5446	-0.1801	-5.7211	5.0279	25.35070	0.8181	1.5853	-3.8157	
ILS	LON	1.6295	-0.2063	-1.3197	2.0743	-1.1067	-2.34030	0.3420	0.4827	1.0405	
	CNP	1.5991	-0.2076	-0.3550	1.5521	-3.4070	-3.37120	0.7056	0.2710	-1.7960	
	UOP	1.5354	0.5446	-0.1800	-5.7211	5.3293	25.35070	0.8181	1.8853	-1.8157	
3SLS	LON	1.2292	0.8744	0.5412	1.1864	-1.3959	-1.56870	0.6660	0.0867	-1.6067	
	CNP	1.2701	0.8483	0.6438	1.1981	-1.6400	-1.52600	0.6677	0.4292	-1.5360	
	UOP	1.2600	1.0866	0.6404	1.1768	-1.0293	-1.17610	0.6684	0.2633	-1.5016	
FIML	LON	1.2588	0.8669	0.5297	1.1861	-1.3966	-1.59970	0.6650	0.0980	-1.6067	
	CNP	1.2480	0.3251	0.9408	1.1990	-1.0151	-1.44560	0.6681	0.4716	-1.4832	
	UOP	1.1875	-0.5734	-0.5883	0.8336	-0.1137	-1.41765	1.0979	0.9303	-0.0066	

Table 2: Performance of estimators using average N = 100, R = 400

		Equation 1			Equation 2			Equation 3			
Estimator	Correlation level	β ₁₃ (1.8)	γ ₁₁ (0.2)	γ ₁₂ (1.2)	β ₂₃ (1.5)	γ ₂₁ (2.5)	γ ₂₃ (2.1)	β ₃₂ (0.9)	γ ₃₂ (0.4)	γ ₃₃ (3.3)	
OLS	LON	1.2744	0.9452	0.5158	1.1867	-1.62010	-1.5670	0.66770	0.1600	0-1.5899	
	CNP	1.2807	1.2591	0.7718	1.1809	-11.24820	-1.1984	0.67240	0.3253	-1.54300	
	UOP	1.2730	1.0590	0.5338	1.1730	-0.99820	-1.4350	0.67240	0.3253	-1.50070	
LIML	LON	21.5492	0.2310	1-1.693	1.3490	-2.17500	-1.9721	0.64460	0.2647	-1.03910	
	CNP	1.5291	-0.2075	0.9593	-0.4286	1.07090	0.4082	-0.54190	0.3864	1.20150	
	UOP	1.5180	0.2179	-1.2211	-2.4245	5.06240	23.4516	0.78770	1.3781	-3.26200	
2SLS	LON	1.5491	0.2310	-1.6931	1.3490	-2.17650	-1.9720	0.64460	0.2647	-1.03870	
	CNP	1.5291	0.9593	-0.4286	1.0709	0.40820	-0.5302	0.38640	1.2015	-1.52780	
	UOP	1.5180	0.2179	-1.2210	-1.4423	`5.06250	23.4510	6.07877	1.3780	-3.26200	
ILS	LON	1.5491	0.2310	-1.6931	1.2490	-2.17510	-1.9721	0.64460	0.2647	-1.03900	
	CNP	1.5292	0.9591	-0.4286	1.0710	0.40850	-0.5303	0.38640	1.2016	-1.52760	
	UOP	1.5180	0.2178	-1.2211	-2.4245	5.06250	23.4515	0.78770	1.3781	-3.26200	
3SLS	LON	1.2547	0.9570	0.3782	1.2081	-1.73220	-1.4841	0.67300	0.0793	-1.46620	
	CNP	1.2219	1.1562	0.8135	1.1800	-1.39200	-1.1753	0.66520	1.2507	-1.46540	
	UOP	1.2810	1.0570	0.5345	1.1716	-1.08710	-1.1631	0.61200	0.4500	-1.50510	
FIML	LON	1.2542	0.9566	0.3784	1.2091	-1.46320	-1.7331	0.67260	0.0795	-1.51510	
	CNP	0.9539	0.6391	-0.3102	1.2251	0.05040	-1.4592	0.65850	0.3414	-1.20720	
	UOP	1.1647	-0.6858	-1.0105	0.6650	-0.04740	-1.3985	1.24620	0.9243	-0.08150	

Table 3: F	Performance of esti	mators using a	average N = 1	00, $R = 600$						
		Equation 1			Equation 2		Equation 3			
Estimator	Correlation level	β ₁₃ (1.8)	γ ₁₁ (0.2)	γ ₁₂ (1.2)	β ₂₃ (1.5)	γ ₂₁ (2.5)	γ ₂₃ (2.1)	β ₃₂ (0.9)	γ ₃₂ (0.4)	γ ₃₃ (3.3)
OLS	LON	1.2860	0.7744	0.6784	1.1650	-14341	-1.9785	0.6660	0.1542	-1.2714
	CNP	1.2915	1.2240	0.8452	1.1788	-1.2065	-1.2334	0.6542	0.3153	-12.7742
	UOP	1.2695	1.2642	0.6645	1.1914	-1.5942	-1.4472	0.6771	0.4366	-1.5071
LIML	LON	1.5293	1.6521	-1.4680	1.1431	-1.6762	-1.3765	0.7158	0.3073	-1.5146
	CNP	1.3892	1.3962	-1.7714	1.1673	-1.0675	-1.9080	0.6367	0.3777	-1.6672
	UOP	1.5921	0.2179	-1.5212	1.3561	-1.3760	-2.5926	0.6882	0.4159	-1.5693
2SLS	LON	1.5293	1.6520	-1.4681	-1.2430	-1.6760	-1.3765	0.7158	0.3072	-1.5146
	CNP	1.3890	1.3962	-1.7710	1.1613	-1.0675	-1.9080	0.6368	0.3777	-1.6670
	UOP	1.5923	0.2179	-1.5210	1.3561	-1.3760	-2.5921	0.6882	0.4159	-1.5692
ILS	LON	1.5294	1.1521	-1.4680	-1.1432	-1.6761	-1.3766	0.7159	0.3073	-1.5146
	CNP	1.3890	1.3962	-1.7713	1.1614	-10.675	-1.9081	0.6368	0.3778	-1.6671
	UOP	1.5921	0.2179	-1.5311	1.3562	-2.36761	-2.5927	0.6882	0.4155	01.5693
3SLS	LON	1.2661	1.2620	0.6026	1.1942	-1.8403	-1.716	0.6801	0.1543	-1.9912
	CNP	1.2914	1.2240	0.8452	1.1787	-1.2066	-1.2335	0.6541	0.3150	-1.7741
	UOP	1.2692	1.2640	0.6843	1.1912	-1.5940	-1.4470	0.6770	0.4371	-1.5070
FIML	LON	1.2660	1.2621	0.6026	1.1940	-1.8401	-1.7102	0.6720	0.1540	-1.9901
	CNP	1.2914	1.2240	0.8520	1.1787	-1.2064	-1.2333	0.6542	0.3153	-1.7741
	UOP	1.1914	-1.5940	-1.4470	0.6771	0.4366	-1.5071	1.2814	0.5393	-0.0103

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Table 4: Performance of estimators using average N = 200, R = 200

		Equation	1		Equation 2			Equation 3		
Estimator	Correlation level	β ₁₃ (1.8)	γ ₁₁ (0.2)	γ ₁₂ (1.2)	β ₂₃ (1.5)	γ ₂₁ (2.5)	γ ₂₃ (2.1)	β ₃₂ (0.9)	γ ₃₂ (0.4)	γ ₃₃ (3.3)
OLS	LON	1.2727	1.0506	0.5946	1.2016	-1.8306	-1.9605	0.6637	0.07650	-1.8643
	CNP	1.2834	1.0068	0.5390	1.1934	-1.4695	-1.3598	0.6661	023830	-1.7775
	UOP	1.2556	1.0571	0.5599	1.1782	-1.3179	-1.4356	0.6687	0.21540	-1.5906
LIML	LON	1.6082	-1.0314	-0.7975	1.4243	-1.5840	-2.5461	0.7962	0.13150	-2.6240
	CNP	1.5536	-0.7173	-0.7303	1.4062	-2.6744	-2.4881	0.7167	0.02057	-1.9184
	UOP	1.7210	-1.9232	0.8270	1.5058	-1.9106	-1.9950	0.4900	0.71770	-1.9487
2SLS	LON	1.6082	-1.0313	-0.7959	1.4242	-1.5841	-2.5461	0.7962	0.13150	-2.6242
	CNP	1.5535	-0.7173	-0.7303	1.4062	-2.0744	-2.4881	0.7167	0.20570	-1.9184
	UOP	1.7210	-1.9232	0.8270	1.5057	-1.9105	-1.9950	0.4700	0.71770	-1.9487
ILS	LON	1.6082	-1.0314	-0.7959	1.4242	-1.5841	-2.5461	0.7962	0.13150	-2.6242
	CNP	1.5535	-0.7173	-0.7305	1.4061	-2.0745	-248810	0.7167	0.20570	-1.9184
	UOP	1.7211	-1.9233	0.8271	1.5057	-1.9106	-1.9950	0.7010	0.71780	-1.9482
3SLS	LON	1.2753	0.9973	0.7320	1.1834	-1.5576	-1.7442	0.6738	0.08900	-1.8642
	CNP	1.2445	0.6317	0.8584	1.1928	-1.1309	-1.4843	0.6829	0.05250	-1.7919
	UOP	1.2551	1.0574	0.8913	1.1747	-1.2150	-1.4172	0.6814	0.53130	-1.7919
FIML	LON	1.2723	0.9943	0.7458	1.2016	-1.8301	-1.9635	0.6694	0.07650	-1.8640
	CNP	1.2908	0.7905	0.9496	1.1864	-1.3412	-1.3597	0.6626	0.50670	-1.4426
	UOP	1.1748	-1.3170	-1.4177	0.6813	0.0587	-1.7916	1.2825	0.93880	-1.0815

Table 5: Performance of estimators using average N = 200, R = 400

					Equation 2			Equation 3			
Estimator	Correlation level	β ₁₃ (1.8)	γ ₁₁ (0.2)	γ ₁₂ (1.2)	$\beta_{23}(1.5)$	γ ₂₁ (2.5)	γ ₂₃ (2.1)	$\beta_{32}(0.9)$	γ ₃₂ (0.4)	γ ₃₃ (3.3)	
OLS	LON	1.27730	0.7043	0.9268	1.1827	-1.4498	-1.3200	0.6637	0.4579	-1.5924	
	CNP	1.26600	0.9280	1.1190	1.1216	-1.0115	-1.2094	0.6541	0.4154	-1.4687	
	UOP	1.25950	0.6891	1.1934	1.2027	-1.3204	-1.1870	0.6649	0.3563	-1.0018	
LIML	LON	1.51870	1.4902	0.7635	3.7819	-1.8406	0.7651	0.6404	0.0685	-1.8759	
	CNP	0.64920	0.8809	1.1260	1.7665	-1.7314	-4.4351	0.7189	0.3516	-2.1434	
	UOP	1.58350	0.2482	-0.7098	1.8379	-3.5292	-3.2115	0.7736	0.2491	-1.7203	
2SLS	LON	1.51870	1.4902	0.7636	3.7819	-1.8408	0.7652	0.6404	0.0698	-1.8763	
	CNP	0.64920	0.8812	1.1263	1.7667	-1.7314	-4.4551	0.7189	0.3516	-2.1435	
	UOP	1.58360	0.2482	-0.7099	1.8378	-2.5290	-3.2116	0.7740	0.2491	-1.7201	
ILS	LON	1.51870	1.4903	0.7635	3.7829	4.8409	0.7652	0.6404	0.0699	-1.8765	
	CNP	0.64920	0.8810	1.1261	1.7668	-1.7314	-4.4549	0.7189	0.3516	-2.1430	
	UOP	1.58360	0.2482	-0.7099	1.8378	-3.5290	-3.2115	0.7740	0.2490	-1.7200	
3SLS	LON	1.26540	0.0891	0.4072	1.1815	-1.7476	-1.5438	0.6645	0.3291	-1.4909	
	CNP	1.26201	0.9279	1.1194	1.2001	-0.6335	-1.3669	0.6650	0.1203	-1.6836	
	UOP	1.26500	0.8079	0.4069	1.1810	-1.7470	-1.5406	0.6641	0.3285	-1.4906	
FIML	LON	1.26500	0.8079	0.4069	1.1812	-1.7470	-1.5406	0.6641	0.3285	1.4906	
	CNP	1.26210	0.9278	1.1194	1.2005	-0.6329	-1.3667	0.6646	0.1202	-1.6837	
	UOP	1.16440	-1.6063	-0.7131	0.6742	0.0371	-1.7175	0.6749	0.5394	-1.5556	

		Equation 1			Equation 2			Equation 3		
Estimator	Correlation level	β ₁₃ (1.8)	γ ₁₁ (0.2)	γ ₁₂ (1.2)	β ₂₃ (1.5)	γ ₂₁ (2.5)	γ ₂₃ (2.1)	β ₃₂ (0.9)	γ ₃₂ (0.4)	γ ₃₃ (3.3)
OLS	LON	1.2381	0.9633	0.6870	1.1615	-1.2390	-1.7460	0.67140	-0.0390	-1.8957
	CNP	1.2426	0.9895	0.2829	1.1680	-1.2037	01.946	0.68390	0.4224	-1.4258
	UOP	1.2449	1.5186	1.2984	1.1972	-1.7290	-1.1683	0.67200	0.3232	-1.8184
LIML	LON	1.3087	0.7384	0.3025	1.5126	-0.7987	-2.7836	0.73360	0.2915	-2.1426
	CNP	0.8928	1.5664	13.5267	1.3024	-2.1758	-1.3090	0.51980	0.9387	-0.0848
	UOP	1.4440	2.2965	-1.3974	1.3280	-1.2032	-2.1286	0.75090	0.3323	-3.2624
2SLS	LON	1.3085	0.7383	0.3024	1.5125	-0.2987	-2.7836	0.73360	0.2915	-2.1426
	CNP	0.8928	1.5664	3.5267	1.3014	-2.1758	-1.3091	0.51980	0.9387	-0.0848
	UOP	1.4440	2.2965	-1.3974	1.3281	-1.3024	-2.1286	0.78090	0.3323	-3.2623
ILS	LON	1.3088	0.7383	0.3025	1.5125	-0.7987	-2.7836	0.73370	0.2915	-2.1426
	CNP	0.8930	1.5660	3.5268	1.3015	-2.1758	-1.3091	0.51980	0.9387	-0.0848
	UOP	1.4443	2.2964	-1.3974	1.3281	-1.3024	-2.1286	0.78090	0.3323	-3.2623
3SLS	LON	0.8395	-1.3294	1.0388	1.1840	1.6564	-1.2007	0.65840	0.1067	-1.7245
	CNP	0.8544	-1.0475	1.7214	1.1659	-1.2335	-1.9890	0.65970	0.1835	-1.2606
	UOP	1.3807	1.8994	0.2365	1.2018	-1.4270	-1.3316	0.67460	0.0454	-1.4465
FIML	LON	1.2382	0.9625	0.6870	1.1616	-1.2394	-1.7460	0.67150	-0.0394	-1.8957
	CNP	0.8501	-1.2316	1.0604	1.1685	1.4590	-1.2257	06.6601	0.3951	-1.1736
	UOP	1.1507	-1.0913	-1.1638	0.6523	-0.0259	-1.5526	1.27320	1.3586	0.45700

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Table 6: Performance of estimators using average N = 200, R = 600

Table 7: Performance of estimators using absolute bias estimates N = 100, R = 200

	Correlation level	Equation	l		Equation	2		Equation 3		
Estimator		β ₁₃ (1.8)	γ ₁₁ (0.8)	γ ₁₂ (1.2)	$\beta_{21}(1.5)$	γ ₂₁ (2.5)	γ ₂₃ (2.1)	β ₃₂ (0.9)	γ ₃₂ (0.4)	γ ₃₃ (3.3)
	LON	0.4942	0.6156	1.0265	0.2321	4.3553	3.6829	0.2488	0.3138	4.8255
OLS	CNP	0.5140	0.8031	0.8280	0.3004	3.9220	3.6639	0.2119	0.2431	4.7746
	UOP	0.5473	0.8212	0.3813	0.3057	3.8813	2.8298	0.2172	0.2037	4.7653
	LON	0.4760	0.1123	1.0491	0.8240	2.5860	1.5318	0.0288	0.0388	4.1323
L2ILS	CNP	0.2717	0.6921	0.6371	0.0894	4.3832	4.0169	0.0306	0.0116	4.2654
	UOP	0.5115	0.8587	0.1646	0.1960	3.2937	3.9453	0.0045	0.0285	4.8924
	LON	0.3931	0.7540	1.2323	0.5193	4.5253	3.7152	0.2066	0.0116	4.6169
3SLS	CNP	0.2891	0.9201	0.6343	0.0772	4.6109	3.3892	0.1214	0.0285	1.7512
	UOP	0.4577	0.7245	0.4926	0.2396	3.7741	2.6789	0.2032	0.0476	4.6823
	LON	0.4790	0.4400	1.0484	0.3390	4.2450	4.3041	0.2575	0.2528	4.6500
FIML	CNP	0.4589	0.9031	0.3428	0.1466	4.8620	4.7718	0.2610	0.1160	5.1917
	UOP	0.5258	0.7432	0.3972	0.1786	3.9094	4.3087	0.1018	0.1769	4.7609

Table 8: Performance of estimators using absolute bias N = 100, R = 400

		Equation 1			Equation 2			Equation 3			
Estimator	Correlation level	β ₁₃ (1.8)	γ ₁₁ (0.8)	$\gamma_{12}(1.2)$	$\beta_{21}(1.5)$	γ ₂₁ (2.5)	γ ₂₃ (2.1)	β ₃₂ (0.9)	γ ₃₂ (0.4)	γ ₃₃ (3.3)	
	LON	0.4529	0.3865	0.3205	0.3155	4.9627	4.2183	0.2339	0.4363	5.6683	
OLS	CNP	0.5603	0.5383	0.5265	0.3437	3.8650	3.6558	0.2106	0.5072	5.2518	
	UOP	0.5541	0.0864	0.1414	0.3130	4.2930	2.8957	0.2177	0.0488	5.0494	
	LON	0.4478	0.0411	1.5486	0.4089	6.9714	5.5440	0.2837	0.0265	5.1042	
L2ILS	CNP	0.6973	8.3829	2.2347	38.3893	412.3580	58.3370	24.6602	277.8090	53.2742	
	UOP	0.5003	0.9963	3.0000	0.2455	6.7980	6.9953	0.1762	0.4848	5.6115	
	LON	0.4478	0.0411	1.5486	0.4089	6.9714	5.5440	0.2837	0.0265	5.1044	
3SLS	CNP	0.4298	0.9873	0.3918	0.1197	6.0204	7.0993	0.2925	0.5875	4.8938	
	UOP	0.6512	1.2496	0.2093	0.2103	5.3973	5.6804	0.2176	1.7704	5.9290	
	LON	0.5047	0.2116	2.5459	2.9107	4.7264	4.9437	0.2330	1.9180	4.4676	
FIML	CNP	0.4575	0.9134	0.1415	0.2008	9.6492	5.4389	0.2132	0.5254	4.5740	
	UOP	0.5979	0.7562	2.6659	0.1767	9.6426	4.2057	0.3266	1.9602	5.2569	

Table 9: The 'best' estimators arranged according to their ranks

Equation 1			Equation 2			Equation 3			All Equations		
LON	CNP	UOP	LON	CNP	UOP	LON	CNP	UOP	LON	CNP	UOP
FIML	FIML	L2ILS	L2ILS	FIML	OLS	L2ILS	3SLS	FIML	L2ILS	FIML	OLS
L2ILS	3SLS	3SLS	3SLS	3SLS	3SLS	3SLS	L2ILS	L2ILS	FIML	3SLS	3SLS
OLS	OLS	0LS	FIML	OLS	L2ILS	FIML	OLS	OLS	3SLS	OLS	L2ILS
3SLS	L2ILS	FIML	OLS	L2ILS	FIML	OLS	FIML	3SLS	OLS	L2ILS	FIML

CONCLUSION

This study has established that L2ILS estimators are best for estimating parameters of data plagued by the Lower Open Interval Negative level of multicollinearity, while OLS performed poorly for this category. In the closed interval which yielded estimates that are closest to the true parameter values, FIML ranks highest while OLS gave a clear lead in estimating parameters of data characterized by the UOP. When compared with some earlier research, this study exhibit smaller biases and suggest that the higher number of equations and parameters may likely reduce the adverse effects of multicollinearity. We further, recommend that only CNP estimates should be used when faced with multicollinearity problems.

Finally, since a Monte Carlo simulation technique was used the scope of generalization is unavoidably limited to all the assumptions made in generating the data sets used.

Areas for further research: The following areas may be explored by other researchers for further contribution to knowledge:

- The effect of inclusion of structural equations with different status of identification (just or over identified)
- The effect of multicollinearity of exogenous variables under more than three levels of correlation coefficients
- The effect of multicollinearity of exogenous variables for models with three structural equations for both upper and Lower triangular

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